

Continuous-Time Markov Chains (CTMC)

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Excerpt from : **CTMC** , Ward Whitt

Definition of a CTMC

- For the continuous-time Markov chain $\{X(t) : t \geq 0\}$ with N states, the Markov property can be written as
- $P[X(s + t) = j \mid X(s) = i, X(u) = x(u), 0 \leq u < s] = P[X(s + t) = j \mid X(s) = i], i, j \in S, 0 \leq t < \infty,$
- and reflects the fact that the future state at time $s+t$ only depends on the current state at time s .

transition probabilities functions

- We consider the special case of stationary transition probabilities functions (sometimes referred to as homogeneous transition probabilities functions), occurring when

$$P[X(s + t) = j \mid X(s) = i] = P[X(t) = j \mid X(0) = i] = P_{ij}(t)$$

for all states i and j and for all times $s > 0$ and $t > 0$;

- the independence of s characterizes the stationarity.

and

$$\mathbf{P}(t) = [P_{ij}(t)]$$

*is called the **transition probability matrix function (TPMF)**.* (a function of time compared to TPM)

Exponential holding time in states of CTMC

- **Proposition:** T_i is exponentially distributed
- **Proof:** By time homogeneity, we assume that the process starts out in state i . For $s \geq 0$ the event $\{T_i > s\}$ is equivalent to the event $\{X(u) = i \text{ for } 0 \leq u \leq s\}$.
- Similarly, for $s, t \geq 0$ the event $\{T_i > s+t\}$ is equivalent to the event $\{X(u) = i \text{ for } 0 \leq u \leq s + t\}$.

Exponential holding time in states of CTMC

- Therefore,

$$P(T_i > s + t | T_i > s)$$

$$= P(X(u) = i \text{ for } 0 \leq u \leq s + t | X(u) = i \text{ for } 0 \leq u \leq s)$$

$$= P(X(u) = i \text{ for } s < u \leq s + t | X(u) = i \text{ for } 0 \leq u \leq s)$$

$$= P(X(u) = i \text{ for } s < u \leq s + t | X(s) = i)$$

$$= P(X(u) = i \text{ for } 0 < u \leq t | X(0) = i)$$

$$= P(T_i > t),$$

Exponential holding time in states of CTMC

- where
 - the second equality follows from the simple fact that $P(A \cap B | A) = P(B | A)$, where we let $A = \{X(u) = i \text{ for } 0 \leq u \leq s\}$ and $B = \{X(u) = i \text{ for } s < u \leq s + t\}$.
 - the third equality follows from the Markov property.
 - the fourth equality follows from time homogeneity.
- Therefore, the distribution of T_i has the memoryless property, which implies that it is exponential.

Chapman-Kolmogorov equations

- **Lemma 1.** (Chapman-Kolmogorov equations) For all $s \geq 0$ and $t \geq 0$, $P_{i,j}(s + t) = \sum_k P_{i,k}(s)P_{k,j}(t)$
- Or in matrix notation $P(s + t) = P(s)P(t)$
- **Proof**
- We can compute $P_{i,j}(s + t)$ by considering all possible places the chain could be at time s .
- We then condition and uncondition, invoking the Markov property to simplify the conditioning; i.e.,

$$P_{i,j}(s + t) = P(X(s + t) = j | X(0) = i)$$

Chapman-Kolmogorov equations

- *Proof (cntd.)*

$$= \sum_k P(X(s+t) = j, X(s) = k | X(0) = i)$$

=

$$\sum_k P(X(s) = k | X(0) = i) P(X(s+t) = j | X(s) = k, X(0) = i)$$

(conditioning on $X(s)=k$)

$$= \sum_k P(X(s) = k | X(0) = i) P(X(s+t) = j | X(s) = k)$$

(Markov property) (uncondition)

$$= \sum_k P_{i,k}(s) P_{k,j}(t) \text{ (stationary transition probabilities)}$$

■

Describing a CTMC

- A CTMC is well specified if we specify:
- (1) its **initial probability distribution** – $p(X(0) = i)$ for all states i
- (2) its **transition probabilities** - $P_{i,j}(t)$ for all states i and j and positive times t .
- Thus we use these two elements to compute the distribution of $X(t)$ for each t ,

$$P(X(t) = j) = \sum_i P(X(0) = i)P_{i,j}(t)$$

Describing a CTMC

- Since the CTMC must be at any time in one of the N states, the analogous of DTMC is, for any state i

$$\sum_{j=1}^N P_{i,j}(t) = 1$$

constructing a CTMC model- **four** approaches(models)

- for all four models:
- the **initial distribution are required** and thus we focus on specifying the model beyond the initial distribution.
- The **four** models are equivalent: you can get to each from any of the others.
- Even though these four approaches are redundant, they are useful because they together give a different more comprehensive view of a CTMC.

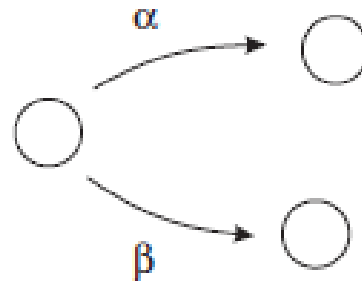
constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

- For the DTMC with transition matrix P (looking at the transition epochs of the CTMC thus $p_{ii}=0$), the transition probabilities of the embedded chain

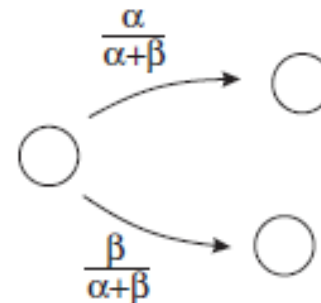
$$\begin{aligned}
 p_{i,j} &= \lim_{\Delta t \rightarrow 0} P\{X_{t+\Delta t} = j \mid X_{t+\Delta t} \neq i, X_t = i\} \\
 &= \lim_{\Delta t \rightarrow 0} \frac{P\{X_{t+\Delta t} = j, X_{t+\Delta t} \neq i \mid X_t = i\}}{P\{X_{t+\Delta t} \neq i \mid X_t = i\}} \\
 &= \begin{cases} \frac{q_{i,j}}{\sum_j q_{i,j}} & i \neq j \\ 0 & i = j \end{cases} \quad \text{cf. } P\{\min(X_1, \dots, X_n) = X_i\} = \frac{\lambda_i}{\lambda_1 + \dots + \lambda_n}, \text{ when } X_i \sim \text{Exp}(\lambda_i)
 \end{aligned}$$

constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

- Markov process, transition **rates** $q_{i,j}$ equilibrium probabilities $\tilde{\pi}_i$



- Embedded Markov chain, transition **probabilities** $p_{i,j}$
- equilibrium probabilities π_i



constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

- For this DTMC the steady state probability vector is π , the unique probability vector satisfying the equation

$$\pi = \pi P \quad (1)$$

- Instead of having each transition take unit time, now we assume that the time required to make a transition from state i has an exponential distribution with **rate** q_i , and thus **mean** $1/q_i$, independent of the history before reaching state i .

constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

- Relating the **steady-state (stationary)** probability vector $\tilde{\pi}$ of the CTMC to steady state probability vector of DTMC π

$$\tilde{\pi}_j = \frac{(\pi_j/q_j)}{\sum_k (\pi_k/q_k)} \quad (2)$$

- Indeed, this first modelling approach corresponds to treating the CTMC as a special case of a semi-Markov process (SMP)
- We assume that there are no one-step transitions from any state to itself in the DTMC (no self-loop); i.e., we assume that $P_{i,i} = 0$ for all i (we look at the chain at transitions)
- this assumption is not critical,(see the third modelling)

constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

Markov processes have no self-loops and their state transitions are characterized by a *generator matrix*, which is analogous to a probability transition matrix. The classification of states have analogous statements for Markov processes where the probability transition matrix is replaced by a generator matrix.

The generator matrix of a Markov process, denoted by Q , has entries that are the *rates* at which the process jumps from state to state. These entries are defined by

$$q_{i,j} = \lim_{\tau \rightarrow 0} \frac{P[X(t+\tau)=j|X(t)=i]}{\tau} \quad i \neq j \quad (3')$$

constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

(We assume that the Markov process is time homogeneous and thus that (3') is independent of t .)

The total rate out of state i is denoted by q_i and equals

$$q_i = \sum_{j \neq i}^{\infty} q_{i,j} \quad (4')$$

The holding time of state i is exponentially distributed with rate q_i .

By definition, we set the diagonal entries of Q equal to minus the total rate,

$$q_{i,i} = -q_i \quad (5')$$

This implies that the row sums of matrix Q equal 0.

constructing a CTMC (model 1 : DTMC with Exponential Transition Times)

- **stationary probabilities** in terms of the generator matrix. Using the results of EMC in SMP (i.e. $\tilde{\pi}_i = \frac{\pi_i^e E[S_i]}{\sum_{j \in S} \pi_j^e E[S_j]}$, $i \in S$,) and multiplying (2) by q_{ji} and summing yields [and using (5') $q_{i,i} = -q_i$ and $\pi_i = \sum_{j \neq i} \pi_j p_{j,i} = \sum_{j \neq i} \pi_j \frac{q_{ji}}{q_j}$ slide 12]

$$\begin{aligned} \sum_{j=0}^{\infty} \tilde{\pi}_j q_{ji} &= \frac{\sum_{j=0}^{\infty} (\pi_j q_{ji} / q_j)}{\sum_k (\pi_j / q_j)} = \frac{\sum_{j \neq i} (\pi_j q_{ji} / q_j) + \pi_i q_{ii} / q_i}{\sum_k (\pi_j / q_j)} \\ &= \frac{\pi_i - \pi_i}{\sum_k (\pi_j / q_j)} = \mathbf{0} \quad [\text{nel (8.65)}] \end{aligned}$$

- Rewriting in matrix form, shows that the stationary probabilities of a Markov process satisfy

$$\pi Q = 0,$$

with the additional normalization requirement that

$$\|\pi\| = 1.$$

constructing a CTMC (model 2 : Transition Rates and ODE's)

- We look at the chain at any time (so we need to define zero-time transition probabilities, $P_{i,i}(0) = 1$ since there is no instant jump from a state)
- let $P(0) = I$, where I is the identity matrix; i.e., we set $P_{i,i}(0) = 1$ for all i and we set $P_{i,j}(0) = 0$ whenever $i \neq j$.

- We **define** $Q \equiv \lim_{h \downarrow 0} \frac{P(h) - I}{h} = \lim_{h \downarrow 0} \frac{P(h) - P(0)}{h}$
= $P'(0+)$ (it is rate)

See Ross prob. Models 9th ed. ch 6 page 378

constructing a CTMC (model 2 : Transition Rates and ODE's)

- Thus the transition rate from state i to state j be defined in terms of the derivatives:

$$Q_{i,j} \equiv \lim_{h \downarrow 0} \frac{P_{i,j}(h) - P_{i,j}(0)}{h} = P'_{i,j}(0+) = \left. \frac{dP_{i,j}(t)}{dt} \right|_{t=0+} \quad (3)$$

$$Q_{i,i} \equiv \lim_{h \downarrow 0} \frac{P_{i,i}(h) - P_{i,i}(0)}{h} = \frac{P_{i,i}(h) - 1}{h} = P'_{i,i}(0+) = \left. \frac{dP_{i,i}(t)}{dt} \right|_{t=0+}$$

- in most treatments of CTMC's instead of above, it is common to assume that

$$P_{i,j}(h) = Q_{i,j}h + o(h) \text{ as } h \downarrow 0 \text{ if } j \neq i \quad (4) \quad \text{and}$$

$$P_{i,i}(h) - 1 = Q_{i,i}h + o(h) \text{ as } h \downarrow 0, \quad (5)$$

constructing a CTMC (model 2 : Transition Rates and ODE's)

- For finite state space, (for infinite state spaces under extra regularity conditions), we have

- $Q_{i,i} = - \sum_{j, j \neq i} Q_{i,j}(t) \quad (6)$

since $P_{i,j}(t)$ sum over j to 1

$$\sum_{j=1}^N P_{i,j}(t) = 1 \text{ so } P_{i,i}(t) + \sum_{j=1, j \neq i}^N P_{i,j}(t) = 1$$

$$\sum_{j=1, j \neq i}^N P_{i,j}(t) = 1 - P_{i,i}(t)$$

Dividing by t and let $t \rightarrow 0$ we obtain (6)

And let

$$-Q_{i,i} = q_i \quad (7) \text{ for all } i,$$

constructing a CTMC (model 2 : Transition Rates and ODE's)

- Same as DTMC model that is specified via a transition probability matrix P , we can specify a CTMC model via the transition-**rate** matrix Q .
- In specifying the transition-rate matrix Q , it suffices to specify the off-diagonal elements $Q_{i,j}$ for $i \neq j$, because the diagonal elements $Q_{i,i}$ are always defined by (6).
- The off-diagonal elements are always nonnegative, whereas the diagonal elements are always negative.
- Each row sum of Q is zero.

constructing a CTMC (model 2 : Transition Rates and ODE's)

- In fact, this approach to CTMC modelling is perhaps best related to modelling with ordinary differential equations,
- We may use Chapman-Kolmogorov equations to find the transition probabilities $P_{i,j}(t)$ from the transition rates $Q_{i,j} \equiv P'_{i,j}(0+)$
- To do this we use the two systems of ordinary differential equations (ODE's) generated by the transition rates namely, Kolmogorov forward and backward ODE's (defined next).

constructing a CTMC (model 2 : Transition Rates and ODE's)

Theorem 1. (Kolmogorov forward and backward ODE's) The transition probabilities satisfy both the Kolmogorov forward differential equations

$$P'_{i,j}(s + t) = \sum_k P_{i,k}(s) Q_{k,j}(t) \quad \text{for all } i, j \quad (9)$$

in matrix notation is the matrix ODE

$$P'(t) = P(t)Q \quad (10)$$

and the Kolmogorov backward differential equations

$$P'_{i,j}(s + t) = \sum_k Q_{i,k}(t) P_{k,j}(t) \quad \text{for all } i, j \quad (11)$$

in matrix notation is the matrix ODE

$$P'(t) = QP(t) \quad (12)$$

constructing a CTMC (model 2 : Transition Rates and ODE's)

- **Proof:** We apply the Chapman-Kolmogorov equations to write

$$P(t + h) = P(t)P(h) ,$$

and then do an asymptotic analysis as $h \downarrow 0$.

- We subtract $P(t)$ from both sides and divide by h , to get

$$\frac{P(t + h) - P(t)}{h} = P(t) \frac{P(h) - I}{h}$$

where I is the identity matrix

constructing a CTMC (model 2 : Transition Rates and ODE's)

- Recalling that $I = P(0)$, we can let $h \downarrow 0$ to get the desired result (10).
- To get the backward equation (12), we start with

$$P(t + h) = P(h + t) = P(h)P(t)$$

and reason in the same way ■

constructing a CTMC (model 2 : Transition Rates and ODE's)M/M/1 Queue

- **Example** (Transient Probabilities for the $M/M/1$ Queue)
- Note that given that the initial state at time 0 was state i ,
- Writing the forward equation for the $M/M/1$ queue yields
- $$\frac{dP_{i,0}(t)}{dt} = \mu P_{i,1}(t) - \lambda P_{i,0}(t),$$
- $$\frac{dP_{i,j}(t)}{dt} = \mu P_{i,j+1}(t) + \lambda P_{i,j-1}(t) - (\lambda + \mu)P_{i,j}(t).$$

constructing a CTMC (model 2 : Transition Rates and ODE's) M/M/1 Queue

- **Example (Cntd.)** The solution to these equations for this case is then given by

$$P_{i,j}(t) = e^{-(\lambda + \mu)t} \left[\rho^{(j-i)/2} I_{j-i}(\alpha t) + \rho^{(j-i-1)/2} I_{j+i+1}(\alpha t) + (1 - \rho)\rho^j \sum_{k=j+i+2}^{\infty} \rho^{-k/2} I_k(\alpha t) \right]$$

- where $\rho = \frac{\lambda}{\mu}$ and $\alpha = 2\mu\sqrt{\rho}$ and

$$I_k(x) = \sum_{m=0}^{\infty} \frac{\left(\frac{x}{2}\right)^{k+2m}}{(k+m)!m!} \quad k \geq -1$$

is the series expansion for the modified Bessel function of the first kind.

constructing a CTMC (model 2 : Transition Rates and ODE's) M/M/1 Queue

- **Example (Cntd.)** It is difficult to have any intuition regarding the solution except for its **limiting, and thus stationary,** values.
- (no need for normalization eq. since initial condition $P(0) = (0, \dots, 0, 1, 0, \dots)$ being in state i at $t=0$ ($p_{ii}(0)=1$) is an extra equation)
- In the **third term (i.e. coefficient $(1 - \rho)\rho^j$)** we see factors corresponding to the stationary distribution.
- it must be $\lim_{t \rightarrow \infty} P_{i,j}(t) = (1 - \rho)\rho^j$ independent of i .
- The solution of transient probabilities suggests that :
- $\lim_{t \rightarrow \infty} e^{-(\lambda + \mu)t} \rho^{(j-i)/2} I_{j-i}(\alpha t) = 0$
- $\lim_{t \rightarrow \infty} e^{-(\lambda + \mu)t} \rho^{(j-i-1)/2} I_{j-i-1}(\alpha t) = 0$
- $\lim_{t \rightarrow \infty} e^{-(\lambda + \mu)t} \sum_{k=j+i+2}^{\infty} \rho^{-k/2} I_k(\alpha t) = 1$

constructing a CTMC (model 2 : Transition Rates and ODE's)

- Equations (10 & 12) are matrix ODE's in t that can be similarly solved as the scalar ODE $f'(t) = qf(t)$ and have matrix exponential solution.
- ($P(0) = I$, the initial condition plays no role) In particular, as a consequence of Theorem 1, and if all entries of Q are bounded, (Q is said to be **uniform**: the name comes from uniformization of CTMC in model 4) we have the following corollary.
- $Q_{i,j} = \infty$ means instantaneous jump from state i upon entering this state

constructing a CTMC (model 2 : Transition Rates and ODE's)

- **Theorem 2.** (matrix exponential representation) The transition function can be expressed as a matrix-exponential function of the rate matrix Q , i.e.,

$$P(t) = e^{Qt} = \sum_{n=0}^{\infty} \frac{Q^n t^n}{n!} \quad (13)$$

This matrix exponential is the unique solution to the two ODE's with **initial condition $P(0) = I$** .

constructing a CTMC (model 2 : Transition Rates and ODE's)

- Proof: If we verify or assume that we can interchange summation and differentiation in (13), we can check that the displayed matrix exponential satisfies the two ODE's

$$\begin{aligned} P'(t) &= \frac{d}{dt} \sum_{n=0}^{\infty} \frac{Q^n t^n}{n!} = \sum_{n=0}^{\infty} \frac{d}{dt} \frac{Q^n t^n}{n!} \\ &= \sum_{n=0}^{\infty} \frac{n Q^n t^{n-1}}{n!} = Q \sum_{n=0}^{\infty} \frac{Q^n t^n}{n!} = Q e^{Qt} \quad \blacksquare \end{aligned}$$

Summary of some Models of Markov Processes

Type of Process		Self-Loops	Holding Time
Semi-Markov Processes		No	Arbitrary
Markov chains	Model 1	Yes	$H_i = 1$
	Model 2	No	Geometric, $E[H_i] = (1 - p_{i,i})^{-1}$
Markov processes	Continuous time	No	Exponential, $E[H_i] = q_i^{-1}$
	Uniformized-Model 1	Yes	$H_i = 1$
	Uniformized-Model 2	No	Geometric, $E[H_i] = q_{\max}/q_i$